





Bayesian methods for Inverse problems of imaging systems

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Seeing outside of a body: Making an image with a camera, a microscope or a telescope

- f(x,y) real scene
- g(x, y) observed image
- ► Forward model: Convolution



$$g(x,y) = \iint f(x',y') h(x-x',y-y') dx' dy' + \epsilon(x,y)$$

h(x,y): Point Spread Function (PSF) of the imaging system

Inverse problem: Image restoration

Given the forward model \mathcal{H} (PSF h(x, y))) and a set of data $g(x_i, y_i), i = 1, \cdots, M$ find f(x, y)

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Making an image with an unfocused camera Forward model: 2D Convolution

$$g(x,y) = \iint f(x',y') h(x-x',y-y') \,\mathrm{d}x' \,\mathrm{d}y' + \epsilon(x,y)$$



Inversion: Image Deconvolution or Restoration





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Making an image of the interior of a body



Forward problem: Knowing the object predict the data Inverse problem: From measured data find the object

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Seeing inside of a body: Computed Tomography

- f(x,y) a section of a real 3D body f(x,y,z)
- $g_{\phi}(r)$ a line of observed radiographe $g_{\phi}(r,z)$
- Forward model:

Line integrals or Radon Transform

$$g_{\phi}(r) = \int_{L_{r,\phi}} f(x,y) \, \mathrm{d}l + \epsilon_{\phi}(r)$$

=
$$\iint f(x,y) \, \delta(r - x \cos \phi - y \sin \phi) \, \mathrm{d}x \, \mathrm{d}y + \epsilon_{\phi}(r)$$

Inverse problem: Image reconstruction

Given the forward model \mathcal{H} (Radon Transform) and a set of data $g_{\phi_i}(r), i = 1, \cdots, M$ find f(x, y)

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Computed Tomography: Radon Transform



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Microwave or ultrasound imaging

Measurs: diffracted wave by the object $g(r_i)$ Unknown quantity: $f(r) = k_0^2(n^2(r) - 1)$ Intermediate quantity : $\phi(r)$

$$egin{aligned} g(m{r}_i) &= \iint_D G_m(m{r}_i,m{r}')\phi(m{r}')\,m{f}(m{r}')\,\,\mathrm{d}m{r}',\,\,m{r}_i\in S \ \phi(m{r}) &= \phi_0(m{r}) + \iint_D G_o(m{r},m{r}')\phi(m{r}')\,m{f}(m{r}')\,\,\mathrm{d}m{r}',\,\,m{r}\in S \end{aligned}$$

Born approximation $(\phi(\mathbf{r}') \simeq \phi_0(\mathbf{r}'))$): $g(\mathbf{r}_i) = \iint_D G_m(\mathbf{r}_i, \mathbf{r}')\phi_0(\mathbf{r}') f(\mathbf{r}') d\mathbf{r}', \ \mathbf{r}_i \in S$

Discretization:

$$\begin{cases} \boldsymbol{g} = \boldsymbol{G}_m \boldsymbol{F} \boldsymbol{\phi} \\ \boldsymbol{\phi} = \boldsymbol{\phi}_0 + \boldsymbol{G}_o \boldsymbol{F} \boldsymbol{\phi} \end{cases} \begin{cases} \boldsymbol{g} = \boldsymbol{H}(\boldsymbol{f}) & \bullet & \bullet \\ \text{with } \boldsymbol{F} = \text{diag}(\boldsymbol{f}) \\ \boldsymbol{H}(\boldsymbol{f}) = \boldsymbol{G}_m \boldsymbol{F} (\boldsymbol{I} - \boldsymbol{G}_o \boldsymbol{F})^{-1} \boldsymbol{\phi}_0 \end{cases} \bullet \bullet \bullet \bullet$$

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Fourier Synthesis in different imaging systems



Forward problem: Given f(x, y) compute $G(\omega_x, \omega_y)$ Inverse problem : Given $G(\omega_x, \omega_y)$ on those algebraic lines, cercles or curves, estimate f(x, y)

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Invers Problems: other examples and applications

- X ray, Gamma ray Computed Tomography (CT)
- Microwave and ultrasound tomography
- Positron emission tomography (PET)
- Magnetic resonance imaging (MRI)
- Photoacoustic imaging
- Radio astronomy
- Geophysical imaging
- Non Destructive Evaluation (NDE) and Testing (NDT) techniques in industry
- Hyperspectral imaging
- Earth observation methods (Radar, SAR, IR, ...)
- Survey and tracking in security systems

3. General formulation of inverse problems and classical methods

General non linear inverse problems:

$$g(s) = [\mathcal{H}f(r)](s) + \epsilon(s), \quad r \in \mathcal{R}, \quad s \in \mathcal{S}$$

Linear models:

$$g(s) = \int f(r) h(r, s) dr + \epsilon(s)$$

If $h(\boldsymbol{r}, \boldsymbol{s}) = h(\boldsymbol{r} - \boldsymbol{s}) \longrightarrow \text{Convolution}.$

Discrete data:

$$g(\boldsymbol{s_i}) = \int h(\boldsymbol{s_i}, \boldsymbol{r}) \, \boldsymbol{f(r)} \, \mathrm{d}\boldsymbol{r} + \epsilon(\boldsymbol{s_i}), \quad i = 1, \cdots, m$$

- ► Inversion: Given the forward model \mathcal{H} and the data $g = \{g(s_i), i = 1, \cdots, m)\}$ estimate f(r)
- Well-posed and Ill-posed problems (Hadamard): existance, uniqueness and stability
- Need for prior information

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Inverse problems: Discretization $g(s_i) = \int h(s_i, r) f(r) dr + \epsilon(s_i), \quad i = 1, \cdots, M$

• f(r) is assumed to be well approximated by

$$f(\boldsymbol{r})\simeq\sum_{j=1}^{N}f_{j}\;b_{j}(\boldsymbol{r})$$

with $\{b_j(\boldsymbol{r})\}$ a basis or any other set of known functions

$$g(\boldsymbol{s}_i) = g_i \simeq \sum_{j=1}^N f_j \int h(\boldsymbol{s}_i, \boldsymbol{r}) \, b_j(\boldsymbol{r}) \, d\boldsymbol{r}, \quad i = 1, \cdots, M$$
$$\boldsymbol{g} = \boldsymbol{H} \boldsymbol{f} + \boldsymbol{\epsilon} \text{ with } H_{ij} = \int h(\boldsymbol{s}_i, \boldsymbol{r}) \, b_j(\boldsymbol{r}) \, d\boldsymbol{r}$$

- H is huge dimensional
- ► LS solution : $\widehat{f} = \arg\min_{f} \{Q(f)\}$ with $Q(f) = \sum_{i} |g_{i} [Hf]_{i}|^{2} = ||g Hf||^{2}$ does not give satisfactory result.

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Inverse problems: Deterministic methods Data matching

Observation model

$$g_i = h_i(f) + \epsilon_i, \quad i = 1, \dots, M \longrightarrow g = H(f) + \epsilon$$

• Misatch between data and output of the model $\Delta(\boldsymbol{g}, \boldsymbol{H}(\boldsymbol{f}))$

$$\widehat{\boldsymbol{f}} = rg\min_{\boldsymbol{f}} \left\{ \Delta(\boldsymbol{g}, \boldsymbol{H}(\boldsymbol{f})) \right\}$$

Examples:

-LS
$$\Delta(g, H(f)) = ||g - H(f)||^2 = \sum_i |g_i - h_i(f)|^2$$

$$-L_p \qquad \Delta(\boldsymbol{g}, \boldsymbol{H}(\boldsymbol{f})) = \|\boldsymbol{g} - \boldsymbol{H}(\boldsymbol{f})\|^p = \sum_i |\boldsymbol{g}_i - h_i(\boldsymbol{f})|^p, \quad 1$$

- KL
$$\Delta(\boldsymbol{g}, \boldsymbol{H}(\boldsymbol{f})) = \sum_{i} g_{i} \ln \frac{g_{i}}{h_{i}(\boldsymbol{f})}$$

 In general, does not give satisfactory results for inverse problems.

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Inverse problems: Regularization theory

Inverse problems = III posed problems \longrightarrow Need for prior information

Functional space (Tikhonov):

$$g = \mathcal{H}(f) + \epsilon \longrightarrow J(f) = ||g - \mathcal{H}(f)||_2^2 + \lambda ||\mathcal{D}f||_2^2$$

Finite dimensional space (Philips & Towmey): $\mathbf{g} = \mathbf{H}(\mathbf{f}) + \boldsymbol{\epsilon}$

- Minimum norme LS (MNLS): $J(\mathbf{f}) = ||\mathbf{g} \mathbf{H}(\mathbf{f})||^2 + \lambda ||\mathbf{f}||^2$ $J(\mathbf{f}) = ||\mathbf{a} - \mathbf{H}(\mathbf{f})||^2 + \lambda ||\mathbf{D}\mathbf{f}||^2$
- Classical regularization:
- More general regularization:

$$J(\boldsymbol{f}) = \mathcal{Q}(\boldsymbol{g} - \boldsymbol{H}(\boldsymbol{f})) + \lambda \Omega(\boldsymbol{D}\boldsymbol{f})$$

or

 $J(\boldsymbol{f}) = \Delta_1(\boldsymbol{g}, \boldsymbol{H}(\boldsymbol{f})) + \lambda \Delta_2(\boldsymbol{f}, \boldsymbol{f}_{\infty})$

- Limitations: • Errors are considered implicitly white and Gaussian
- Limited prior information on the solution
- Lack of tools for the determination of the hyperparameters

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Deconvolution example



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Image restoration example

سلام صبح به خير سلام صبح به خير امیدوارم که امیدوارم که روز خوبي داشته باشيد روز خوبی داشته باشید چنین گفت فردوسی راستگوی چنین گفت فردوسی راستگوی ز گهواره تا گور دانش بجوی ز گهواره تا گور دانش بجوی original fPSF h Blurred & noisy qسلام صبح به خبر اميدوارم که روز حوس داشته باشيد چنین گفت فردوسی راستگوی وبين كعب فرءوسيه راستكود ر گهراره با گور دانش بجوی ر گهواره تا گور دانش بجوی ر گهواره با گور داسل بحوی Quadratic reg. \hat{f} L_1 reg. \hat{f} Wiener \hat{f}

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Inversion: Probabilistic methods

Taking account of errors and uncertainties $\longrightarrow \mathsf{Probability}$ theory

- Maximum Likelihood (ML)
- Minimum Inaccuracy (MI)
- Probability Distribution Matching (PDM)
- Maximum Entropy (ME) and Information Theory (IT)
- ► Bayesian Inference (BAYES)

Advantages:

- Explicit account of the errors and noise
- A large class of priors via explicit or implicit modeling
- A coherent approach to combine information content of the data and priors

Limitations:

Practical implementation and cost of calculation

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4. Bayesian inference for inverse problems

$$\mathcal{M}: \quad \boldsymbol{g} = \boldsymbol{H}\boldsymbol{f} + \boldsymbol{\epsilon}$$

 \blacktriangleright Observation model $\mathcal{M}+\mathsf{Hypothesis}$ on the noise $\epsilon\longrightarrow$

$$p(\boldsymbol{g}|\boldsymbol{f};\mathcal{M}) = p_{\boldsymbol{\epsilon}}(\boldsymbol{g} - \boldsymbol{H}\boldsymbol{f})$$

• A priori information $p(\mathbf{f}|\mathcal{M})$

► Bayes : $p(\boldsymbol{f}|\boldsymbol{g};\mathcal{M}) = \frac{p(\boldsymbol{g}|\boldsymbol{f};\mathcal{M}) p(\boldsymbol{f}|\mathcal{M})}{p(\boldsymbol{g}|\mathcal{M})}$

Link with regularization :

Maximum A Posteriori (MAP) :

$$\widehat{\boldsymbol{f}} = \arg \max_{\boldsymbol{f}} \{ p(\boldsymbol{f}|\boldsymbol{g}) \} = \arg \max_{\boldsymbol{f}} \{ p(\boldsymbol{g}|\boldsymbol{f}) \ p(\boldsymbol{f}) \}$$
$$= \arg \min_{\boldsymbol{f}} \{ -\ln p(\boldsymbol{g}|\boldsymbol{f}) - \ln p(\boldsymbol{f}) \}$$

with $Q(\boldsymbol{g}, \boldsymbol{H}\boldsymbol{f}) = -\ln p(\boldsymbol{g}|\boldsymbol{f})$ and $\lambda \Omega(\boldsymbol{f}) = -\ln p(\boldsymbol{f})$

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Case of linear models and Gaussian priors $g = Hf + \epsilon$

- Hypothesis on the noise: $\boldsymbol{\epsilon} \sim \mathcal{N}(0,\sigma_{\epsilon}^{2}\boldsymbol{I})$

$$p(\boldsymbol{g}|\boldsymbol{f}) \propto \exp\left[-\frac{1}{2\sigma^2}\|\boldsymbol{g} - \boldsymbol{H}\boldsymbol{f}\|^2 + \text{Hypothesis on } \boldsymbol{f} : \boldsymbol{f} \sim \mathcal{N}(0, \sigma_f^2 \boldsymbol{I})^\epsilon\right]$$

$$p(\boldsymbol{f}) \propto \exp\left[-\frac{1}{2\sigma_f^2} \|\boldsymbol{f}\|^2\right]$$

A posteriori:

$$p(\boldsymbol{f}|\boldsymbol{g}) \propto \exp\left[-\frac{1}{2\sigma_{\epsilon}^{2}}\left(\|\boldsymbol{g} - \boldsymbol{H}\boldsymbol{f}\|^{2} + \frac{\sigma_{\epsilon}^{2}}{\sigma_{f}^{2}}\|\boldsymbol{f}\|^{2}\right)\right]$$

$$\blacktriangleright \mathsf{MAP}: \quad \widehat{\boldsymbol{f}} = \arg\max_{\boldsymbol{f}} \{p(\boldsymbol{f}|\boldsymbol{g})\} = \arg\min_{\boldsymbol{f}} \{J(\boldsymbol{f})\}$$
with
$$J(\boldsymbol{f}) = \|\boldsymbol{g} - \boldsymbol{H}\boldsymbol{f}\|^{2} + \lambda \|\boldsymbol{f}\|^{2}, \qquad \lambda = \frac{\sigma_{\epsilon}^{2}}{\sigma_{f}^{2}}$$

Advantage : characterization of the solution

$$\boldsymbol{f}|\boldsymbol{g} \sim \mathcal{N}(\widehat{\boldsymbol{f}}, \widehat{\boldsymbol{P}}) \text{ with } \widehat{\boldsymbol{f}} = \left(\boldsymbol{H}^t \boldsymbol{H} + \lambda \boldsymbol{I}\right)^{-1} \boldsymbol{H}^t \boldsymbol{g}, \ \widehat{\boldsymbol{P}} = \sigma_{\epsilon}^2 \left(\boldsymbol{H}^t \boldsymbol{H} + \lambda \boldsymbol{I}\right)^{-1}$$

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MAP estimation with other priors:

$$\widehat{\boldsymbol{f}} = \arg\min_{\boldsymbol{f}} \left\{ J(\boldsymbol{f}) \right\} \text{ with } J(\boldsymbol{f}) = \|\boldsymbol{g} - \boldsymbol{H}\boldsymbol{f}\|^2 + \lambda \Omega(\boldsymbol{f})$$

Separable priors:

Gaussian: $p(f_i) \propto \exp\left[-\alpha |f_i|^2\right] \longrightarrow \Omega(\mathbf{f}) = \|\mathbf{f}\|^2 = \alpha \sum_i |f_i|^2$ Gamma: $p(f_i) \propto f_i^{\alpha} \exp\left[-\beta f_i\right] \longrightarrow \Omega(\mathbf{f}) = \alpha \sum_i \ln f_i + \beta f_i$ Beta: $p(f_j) \propto f_j^{\alpha} (1 - f_j)^{\beta} \longrightarrow \Omega(\mathbf{f}) = \alpha \sum_j \ln f_j + \beta \sum_j \ln(1 - f_j)$ Generalized Gaussian: $p(f_i) \propto \exp\left[-\alpha |f_i|^p\right], \quad 1$ Markovian models: $p(f_j|\boldsymbol{f}) \propto \exp\left[-\alpha \sum_{i \in N_j} \phi(f_j, f_i)\right] \longrightarrow \Omega(\boldsymbol{f}) = \alpha \sum_j \sum_{i \in N_j} \phi(f_j, f_i),$

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MAP estimation and Compressed Sensing

$$\left\{ egin{array}{ll} oldsymbol{g} = oldsymbol{H}oldsymbol{f} + oldsymbol{\epsilon} \ oldsymbol{f} = oldsymbol{W}oldsymbol{z} \end{array}
ight.$$

▶ W a code book matrix, *z* coefficients

Gaussian:

$$p(\boldsymbol{z}) = \mathcal{N}(0, \sigma_z^2 \boldsymbol{I}) \propto \exp\left[-\frac{1}{2\sigma_z^2} \sum_j |z_j|^2\right]$$
$$J(\boldsymbol{z}) = -\ln p(\boldsymbol{z}|\boldsymbol{g}) = \|\boldsymbol{g} - \boldsymbol{H} \boldsymbol{W} \boldsymbol{z}\|^2 + \lambda \sum_j |z_j|^2$$

• Generalized Gaussian (sparsity, $\beta = 1$):

$$p(\boldsymbol{z}) \propto \exp\left[-\lambda \sum_{j} |\boldsymbol{z}_{j}|^{\beta}\right]$$
$$J(\boldsymbol{z}) = -\ln p(\boldsymbol{z}|\boldsymbol{g}) = \|\boldsymbol{g} - \boldsymbol{H}\boldsymbol{W}\boldsymbol{z}\|^{2} + \lambda \sum_{j} |\boldsymbol{z}_{j}|^{\beta}$$

$$\blacktriangleright \mathbf{z} = \operatorname{arg\,min}_{\mathbf{z}} \{J(\mathbf{z})\} \longrightarrow \widehat{\mathbf{f}} = \mathbf{W} \widehat{\mathbf{z}}$$

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Bayesian Estimation: Two simple priors

Example 1: Linear Gaussian case:

$$\begin{cases} p(\boldsymbol{g}|\boldsymbol{f}, \theta_1) = \mathcal{N}(\boldsymbol{H}\boldsymbol{f}, \theta_1 \boldsymbol{I}) \\ p(\boldsymbol{f}|\theta_2) = \mathcal{N}(0, \theta_2 \boldsymbol{I}) \end{cases} \longrightarrow p(\boldsymbol{f}|\boldsymbol{g}, \boldsymbol{\theta}) = \mathcal{N}(\widehat{\boldsymbol{f}}, \widehat{\boldsymbol{P}})$$

with

$$\begin{cases} \widehat{\boldsymbol{f}} = (\boldsymbol{H}'\boldsymbol{H} + \lambda \boldsymbol{I})^{-1}\boldsymbol{H}'\boldsymbol{g} \\ \widehat{\boldsymbol{P}} = \theta_1(\boldsymbol{H}'\boldsymbol{H} + \lambda \boldsymbol{I})^{-1}, \quad \lambda = \frac{\theta_1}{\theta_2} \end{cases}$$
$$\widehat{\boldsymbol{f}} = \arg\min_{\boldsymbol{f}} \left\{ J(\boldsymbol{f}) \right\} \text{ with } J(\boldsymbol{f}) = \|\boldsymbol{g} - \boldsymbol{H}\boldsymbol{f}\|_2^2 + \lambda \|\boldsymbol{f}\|_2^2$$

Example 2: Double Exponential prior & MAP:

$$\widehat{m{f}} = rg\min_{m{f}} \left\{ J(m{f})
ight\}$$
 with $J(m{f}) = \|m{g} - m{H}m{f}\|_2^2 + \lambda \|m{f}\|_1^2$

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Deconvolution example



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Main advantages of the Bayesian approach

- MAP = Regularization
- Posterior mean ? Marginal MAP ?
- More information in the posterior law than only its mode or its mean
- Meaning and tools for estimating hyper parameters
- Meaning and tools for model selection
- More specific and specialized priors, particularly through the hidden variables
- More computational tools:

...

- Expectation-Maximization for computing the maximum likelihood parameters
- MCMC for posterior exploration
- Variational Bayes for analytical computation of the posterior marginals
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Two main steps in the Bayesian approach

- Prior modeling
 - Separable:
 - Gaussian, Gamma,

Sparsity enforcing: Generalized Gaussian, mixture of Gaussians, mixture of Gammas, ...

Markovian:

Gauss-Markov, GGM, ...

 Markovian with hidden variables (contours, region labels)

Choice of the estimator and computational aspects

- MAP, Posterior mean, Marginal MAP
- MAP needs optimization algorithms
- Posterior mean needs integration methods
- Marginal MAP and Hyperparameter estimation need integration and optimization
- Approximations:
 - Gaussian approximation (Laplace)
 - Numerical exploration MCMC
 - Variational Bayes (Separable approximation)

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Sparsity enforcing prior models

► Sparse signals: Direct sparsity



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Sparsity enforcing prior models

- Simple heavy tailed models:
 - Generalized Gaussian (particular case: Double Exponential)
 - Student-t (particular case: Cauchy)
 - Elastic net
 - Symmetric Weibull, Symmetric Rayleigh
 - Generalized hyperbolic
- Hierarchical mixture models:
 - Mixture of Gaussians
 - Bernoulli-Gaussian
 - Mixture of Gammas
 - Bernoulli-Gamma
 - Mixture of Dirichlet
 - Bernoulli-Multinomial

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Simple heavy tailed models

• Generalized Gaussian ((particular case: Double Exponential)

$$p(\mathbf{f}|\boldsymbol{\gamma},\boldsymbol{\beta}) = \prod_{j} \mathcal{GG}(\mathbf{f}_{j}|\boldsymbol{\gamma},\boldsymbol{\beta}) \propto \exp\left[-\gamma \sum_{j} |\mathbf{f}_{j}|^{\boldsymbol{\beta}}\right]$$

- $\beta = 1$ Double exponential or Laplace. $0 < \beta \leq 1$ are of great interest for sparsity enforcing.
- Student-t ((particular case: Cauchy models)

$$p(\mathbf{f}|\nu) = \prod_{j} \mathcal{S}t(\mathbf{f}_{j}|\nu) \propto \exp\left[-\frac{\nu+1}{2}\sum_{j}\log\left(1+\mathbf{f}_{j}^{2}/\nu\right)\right]$$

Cauchy model is obtained when $\nu = 1$.

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Mixture models

• Mixture of two Gaussians (MoG2) model

$$p(\mathbf{f}|\lambda, v_1, v_0) = \prod_j \left(\lambda \mathcal{N}(\mathbf{f}_j|0, v_1) + (1-\lambda)\mathcal{N}(\mathbf{f}_j|0, v_0)\right)$$

• Bernoulli-Gaussian (BG) model

$$p(\boldsymbol{f}|\lambda, v) = \prod_{j} p(\boldsymbol{f}_{j}) = \prod_{j} \left(\lambda \mathcal{N}(\boldsymbol{f}_{j}|0, v) + (1-\lambda)\delta(\boldsymbol{f}_{j}) \right)$$

• Mixture of Gammas

$$p(\boldsymbol{f}|\lambda, v_1, v_0) = \prod_j \left(\lambda \mathcal{G}(\boldsymbol{f}_j | \alpha_1, \beta_1) + (1 - \lambda) \mathcal{G}(\boldsymbol{f}_j | \alpha_2, \beta_2) \right)$$

• Bernoulli-Gamma model

$$p(\mathbf{f}|\lambda, \alpha, \beta) = \prod_{j} \left[\lambda \mathcal{G}(\mathbf{f}_{j}|\alpha, \beta) + (1-\lambda)\delta(\mathbf{f}_{j}) \right]$$

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6. Full Bayesian approach $\mathcal{M}: \quad g = Hf + \epsilon$

- ► Forward & errors model: $\longrightarrow p(\boldsymbol{g}|\boldsymbol{f}, \boldsymbol{\theta}_1; \mathcal{M})$
- Prior models $\longrightarrow p(\mathbf{f}|\boldsymbol{\theta}_2; \mathcal{M})$
- Hyperparameters $\boldsymbol{\theta} = (\boldsymbol{\theta}_1, \boldsymbol{\theta}_2) \longrightarrow p(\boldsymbol{\theta}|\mathcal{M})$
- ► Bayes: $\longrightarrow p(\mathbf{f}, \mathbf{\theta} | \mathbf{g}; \mathcal{M}) = \frac{p(\mathbf{g} | \mathbf{f}, \mathbf{\theta}; \mathcal{M}) p(\mathbf{f} | \mathbf{\theta}; \mathcal{M}) p(\mathbf{\theta} | \mathcal{M})}{p(\mathbf{g} | \mathcal{M})}$
- ► Joint MAP: $(\widehat{f}, \widehat{\theta}) = \arg \max_{(f, \theta)} \{ p(f, \theta | g; M) \}$
- Marginalization 1:

$$p(\boldsymbol{f}|\boldsymbol{g};\mathcal{M}) = \iint p(\boldsymbol{f},\boldsymbol{\theta}|\boldsymbol{g};\mathcal{M}) \; \mathrm{d}\boldsymbol{\theta} \longrightarrow \widehat{\boldsymbol{f}}$$

Marginalization 2:

$$p(\boldsymbol{\theta}|\boldsymbol{g};\mathcal{M}) = \iint p(\boldsymbol{f},\boldsymbol{\theta}|\boldsymbol{g};\mathcal{M}) \, \mathrm{d}\boldsymbol{f} \longrightarrow \widehat{\boldsymbol{\theta}}$$

• Approximate $p(\mathbf{f}, \boldsymbol{\theta} | \mathbf{g}; \mathcal{M})$ by a separable one $q(\mathbf{f}, \boldsymbol{\theta}) = q_1(\mathbf{f})q_2(\boldsymbol{\theta})$ and then use them separately to find $\hat{\mathbf{f}}$ and $\hat{\boldsymbol{\theta}}$.

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Summary of Bayesian estimation 1

Simple Bayesian Model and Estimation



Full Bayesian Model and Hyperparameter Estimation



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Summary of Bayesian estimation 2

Marginalization 1

$$p(\boldsymbol{f}, \boldsymbol{\theta} | \boldsymbol{g}) \longrightarrow p(\boldsymbol{\theta} | \boldsymbol{g}) \longrightarrow \widehat{\boldsymbol{f}}$$

Joint Posterior Marginalize over θ

Marginalization 2

$$p(\boldsymbol{f},\boldsymbol{\theta}|\boldsymbol{g}) \longrightarrow p(\boldsymbol{\theta}|\boldsymbol{g}) \longrightarrow \widehat{\boldsymbol{\theta}} \longrightarrow p(\boldsymbol{f}|\widehat{\boldsymbol{\theta}},\boldsymbol{g}) \longrightarrow \widehat{\boldsymbol{f}}$$

Joint Posterior Marginalize over f

Variational Bayesian Approximation

$$\begin{array}{c} \hline p(\boldsymbol{f}, \boldsymbol{\theta} | \boldsymbol{g}) \end{array} \longrightarrow \begin{array}{c} \begin{array}{c} \mathsf{Variational} \\ \mathsf{Bayesian} \\ \mathsf{Approximation} \end{array} \xrightarrow{} q_1(\boldsymbol{f}) \longrightarrow \widehat{\boldsymbol{f}} \\ \longrightarrow q_2(\boldsymbol{\theta}) \longrightarrow \widehat{\boldsymbol{\theta}} \end{array}$$

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Variational Bayesian Approximation

- ► Full Bayesian: $p(f, \theta|g) \propto p(g|f, \theta_1) p(f|\theta_2) p(\theta)$
- Approximate $p(\mathbf{f}, \boldsymbol{\theta}|\mathbf{g})$ by $q(\mathbf{f}, \boldsymbol{\theta}|\mathbf{g}) = q_1(\mathbf{f}|\mathbf{g}) q_2(\boldsymbol{\theta}|\mathbf{g})$ and then continue computations.
- Criterion $\mathsf{KL}(q(\boldsymbol{f}, \boldsymbol{\theta}|\boldsymbol{g}) : p(\boldsymbol{f}, \boldsymbol{\theta}|\boldsymbol{g}))$

•
$$\mathsf{KL}(q:p) = \int \int q \ln q/p = \int \int q_1 q_2 \ln \frac{q_1 q_2}{p}$$

• Iterative algorithm $q_1 \longrightarrow q_2 \longrightarrow q_1 \longrightarrow q_2, \cdots$

$$\begin{cases} \widehat{q}_{1}(\boldsymbol{f}) \propto \exp \begin{bmatrix} \langle \ln p(\boldsymbol{g}, \boldsymbol{f}, \boldsymbol{\theta}; \mathcal{M}) \rangle_{\widehat{q}_{2}(\boldsymbol{\theta})} \end{bmatrix} \\ \widehat{q}_{2}(\boldsymbol{\theta}) \propto \exp \begin{bmatrix} \langle \ln p(\boldsymbol{g}, \boldsymbol{f}, \boldsymbol{\theta}; \mathcal{M}) \rangle_{\widehat{q}_{1}(\boldsymbol{f})} \end{bmatrix} \end{cases}$$

$$\xrightarrow{p(\boldsymbol{f}, \boldsymbol{\theta}|\boldsymbol{g})} \longrightarrow \begin{cases} \text{Variational} \\ \text{Bayesian} \\ \text{Approximation} \end{cases} \xrightarrow{\rightarrow q_{2}(\boldsymbol{\theta}) \longrightarrow \widehat{\boldsymbol{\theta}}}$$

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7. Hierarchical models and hidden variables

All the mixture models and some of simple models can be modeled via hidden variables z.

$$p(f) = \sum_{k=1}^{K} \alpha_k p_k(f) \longrightarrow \begin{cases} p(f|\mathbf{z}=k) = p_k(f), \\ P(\mathbf{z}=k) = \alpha_k, \quad \sum_k \alpha_k = 1 \end{cases}$$

Example 2: Student-t model

$$St(f|\nu) \propto \exp\left[-\frac{\nu+1}{2}\log\left(1+f^2/\nu\right)\right]$$

Infinite mixture

$$\mathcal{S}t(f|\nu) \propto = \int_0^\infty \mathcal{N}(f|, 0, 1/z) \mathcal{G}(z|\alpha, \beta) \, \mathrm{d}z, \quad \text{with } \alpha = \beta = \nu/2$$

$$\begin{cases} p(\boldsymbol{f}|\boldsymbol{z}) &= \prod_{j} p(f_{j}|\boldsymbol{z}_{j}) = \prod_{j} \mathcal{N}(f_{j}|0, 1/\boldsymbol{z}_{j}) \propto \exp\left[-\frac{1}{2} \sum_{j} \boldsymbol{z}_{j} f_{j}^{2}\right] \\ p(\boldsymbol{z}|\alpha, \beta) &= \prod_{j} \mathcal{G}(\boldsymbol{z}_{j}|\alpha, \beta) \propto \prod_{j} \boldsymbol{z}_{j}^{(\alpha-1)} \exp\left[-\beta \boldsymbol{z}_{j}\right] \\ &\propto \exp\left[\sum_{j} (\alpha-1) \ln \boldsymbol{z}_{j} - \beta \boldsymbol{z}_{j}\right] \end{cases}$$

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Summary of Bayesian estimation 3

• Full Bayesian Hierarchical Model with Hyperparameter Estimation

 $\downarrow oldsymbol{lpha},oldsymbol{eta},oldsymbol{\gamma}$



• Full Bayesian Hierarchical Model and Variational Approximation $\downarrow lpha, eta, \gamma$



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8. Bayesian Computation and Algorithms for Hierarchical models

- Often, the expression of $p(\mathbf{f}, \mathbf{z}, \boldsymbol{\theta} | \mathbf{g})$ is complex.
- Its optimization (for Joint MAP) or its marginalization or integration (for Marginal MAP or PM) is not easy
- Two main techniques: MCMC and Variational Bayesian Approximation (VBA)
- ► MCMC:

Needs the expressions of the conditionals $p(\boldsymbol{f}|\boldsymbol{z}, \boldsymbol{\theta}, \boldsymbol{g}), \; p(\boldsymbol{z}|\boldsymbol{f}, \boldsymbol{\theta}, \boldsymbol{g}), \; \text{and} \; p(\boldsymbol{\theta}|\boldsymbol{f}, \boldsymbol{z}, \boldsymbol{g})$

▶ VBA: Approximate $p(f, z, \theta|g)$ by a separable one

$$q(\boldsymbol{f}, \boldsymbol{z}, \boldsymbol{\theta} | \boldsymbol{g}) = q_1(\boldsymbol{f}) q_2(\boldsymbol{z}) q_3(\boldsymbol{\theta})$$

and do any computations with these separable ones.

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Which images I am looking for?



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Which image I am looking for?



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9. Gauss-Markov-Potts prior models for images



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Four different cases

To each pixel of the image is associated 2 variables $f(\mathbf{r})$ and $z(\mathbf{r})$

- ► f|z Gaussian iid, z iid : Mixture of Gaussians
- ► f|z Gauss-Markov, z iid : Mixture of Gauss-Markov
- ► f|z Gaussian iid, z Potts-Markov : Mixture of Independent Gaussians (MIG with Hidden Potts)
- f|z Markov, z Potts-Markov : Mixture of Gauss-Markov (MGM with hidden Potts)



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Application of CT in NDT

Reconstruction from only 2 projections





$$g_1(x) = \int f(x,y) \,\mathrm{d}y, \qquad g_2(y) = \int f(x,y) \,\mathrm{d}x$$

- Given the marginals $g_1(x)$ and $g_2(y)$ find the joint distribution f(x, y).
- ► Infinite number of solutions : $f(x, y) = g_1(x) g_2(y) \Omega(x, y)$ $\Omega(x, y)$ is a Copula:

$$\int \Omega(x,y) \, \mathrm{d}x = 1 \quad \text{and} \quad \int \Omega(x,y) \, \mathrm{d}y = 1 \quad \text{and} \quad \int \Omega(x,y) \, \mathrm{d}y = 1 \quad \text{and} \quad \mathbb{R}$$

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Application in CT



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Proposed algorithm

 $p(\boldsymbol{f}, \boldsymbol{z}, \boldsymbol{\theta} | \boldsymbol{g}) \propto p(\boldsymbol{g} | \boldsymbol{f}, \boldsymbol{z}, \boldsymbol{\theta}) p(\boldsymbol{f} | \boldsymbol{z}, \boldsymbol{\theta}) p(\boldsymbol{\theta})$

General scheme:

$$\widehat{\boldsymbol{f}} \sim p(\boldsymbol{f} | \widehat{\boldsymbol{z}}, \widehat{\boldsymbol{\theta}}, \boldsymbol{g}) \longrightarrow \widehat{\boldsymbol{z}} \sim p(\boldsymbol{z} | \widehat{\boldsymbol{f}}, \widehat{\boldsymbol{\theta}}, \boldsymbol{g}) \longrightarrow \widehat{\boldsymbol{\theta}} \sim (\boldsymbol{\theta} | \widehat{\boldsymbol{f}}, \widehat{\boldsymbol{z}}, \boldsymbol{g})$$

Iterative algorithme:

- Estimate \boldsymbol{f} using $p(\boldsymbol{f}|\hat{\boldsymbol{z}}, \hat{\boldsymbol{\theta}}, \boldsymbol{g}) \propto p(\boldsymbol{g}|\boldsymbol{f}, \boldsymbol{\theta}) p(\boldsymbol{f}|\hat{\boldsymbol{z}}, \hat{\boldsymbol{\theta}})$ Needs optimisation of a quadratic criterion.
- Estimate \boldsymbol{z} using $p(\boldsymbol{z}|\widehat{\boldsymbol{f}}, \widehat{\boldsymbol{\theta}}, \boldsymbol{g}) \propto p(\boldsymbol{g}|\widehat{\boldsymbol{f}}, \widehat{\boldsymbol{z}}, \widehat{\boldsymbol{\theta}}) p(\boldsymbol{z})$ Needs sampling of a Potts Markov field.
- ► Estimate θ using $p(\theta | \hat{f}, \hat{z}, g) \propto p(g | \hat{f}, \sigma_{\epsilon}^2 I) p(\hat{f} | \hat{z}, (m_k, v_k)) p(\theta)$ Conjugate priors \longrightarrow analytical expressions.

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Results



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Application in Microwave imaging

(Linearized Spectral method: Fourier Synthesis)

$$g(\boldsymbol{\omega}) = \int f(\boldsymbol{r}) \exp\left[-j(\boldsymbol{\omega}.\boldsymbol{r})\right] \, \mathrm{d}\boldsymbol{r} + \epsilon(\boldsymbol{\omega})$$

$$g(u,v) = \iint f(x,y) \exp\left[-j(ux+vy)\right] \, \mathrm{d}x \, \mathrm{d}y + \epsilon(u,v)$$

 $g = Hf + \epsilon$



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Color (Multi-spectral) image deconvolution



Observation model : $\boldsymbol{g}_i = \boldsymbol{H} \boldsymbol{f}_i + \boldsymbol{\epsilon}_i, \quad i = 1, 2, 3$





Same segmentation z for the three components f_i , i = 1, 2, 3 and 2, 3 A. Mohammad-Djafari, Bayesian methods for Inverse problems of imaging systems, Tehran Univ., EECE Dept., December 8, 2014, 47/68

Images fusion and joint segmentation (with O. Féron)

$$\begin{cases} g_{i}(\boldsymbol{r}) = f_{i}(\boldsymbol{r}) + \epsilon_{i}(\boldsymbol{r}) \\ p(f_{i}(\boldsymbol{r})|z(\boldsymbol{r}) = k) = \mathcal{N}(m_{ik}, \sigma_{ik}^{2}) \\ p(\underline{\boldsymbol{f}}|\boldsymbol{z}) = \prod_{i} p(f_{i}|\boldsymbol{z}) \\ p(\boldsymbol{z}) \propto \exp\left[\gamma \sum_{\boldsymbol{r} \in \mathcal{R}} \sum_{\boldsymbol{r}' \in \mathcal{V}(\boldsymbol{r})} \delta(z(\boldsymbol{r}) - z(\boldsymbol{r}'))\right] \end{cases}$$



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Data fusion in medical imaging (with O. Féron)

$$\begin{cases} g_{i}(\boldsymbol{r}) = f_{i}(\boldsymbol{r}) + \epsilon_{i}(\boldsymbol{r}) \\ p(f_{i}(\boldsymbol{r})|z(\boldsymbol{r}) = k) = \mathcal{N}(m_{ik}, \sigma_{ik}^{2}) \\ p(\underline{\boldsymbol{f}}|\boldsymbol{z}) = \prod_{i} p(\boldsymbol{f}_{i}|\boldsymbol{z}) \\ p(\boldsymbol{z}) \propto \exp\left[\gamma \sum_{\boldsymbol{r} \in \mathcal{R}} \sum_{\boldsymbol{r}' \in \mathcal{V}(\boldsymbol{r})} \delta(z(\boldsymbol{r}) - z(\boldsymbol{r}'))\right] \end{cases}$$



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Super-Resolution (with F. Humblot)

$$\begin{aligned} \boldsymbol{g_i(\boldsymbol{r})} &= [\mathcal{DMB}\boldsymbol{f_i}(\boldsymbol{r}) + \epsilon_i(\boldsymbol{r}) \\ p(f_i(\boldsymbol{r})|\boldsymbol{z}(\boldsymbol{r}) = \boldsymbol{k}) &= \mathcal{N}(m_{ik}, \sigma_{i|k}^2) \\ p(\underline{\boldsymbol{f}}|\boldsymbol{z}) &= \prod_i p(\boldsymbol{f_i}|\boldsymbol{z}) \\ p(\boldsymbol{z}) &\propto \exp\left[\gamma \sum_{\boldsymbol{r} \in \mathcal{R}} \sum_{\boldsymbol{r}' \in \mathcal{V}(\boldsymbol{r})} \delta(\boldsymbol{z}(\boldsymbol{r}) - \boldsymbol{z}(\boldsymbol{r}'))\right] \end{aligned}$$



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Joint segmentation of hyper-spectral images

(with N. Bali & A. Mohammadpour)

$$\begin{array}{l} \begin{array}{l} \begin{array}{l} g_{i}(\boldsymbol{r}) = f_{i}(\boldsymbol{r}) + \epsilon_{i}(\boldsymbol{r}) \\ p(f_{i}(\boldsymbol{r})|z(\boldsymbol{r}) = k) = \mathcal{N}(m_{ik}, \sigma_{i\,k}^{2}), \quad k = 1, \cdots, K \\ p(\underline{\boldsymbol{f}}|\boldsymbol{z}) = \prod_{i} p(f_{i}|\boldsymbol{z}) \\ p(\boldsymbol{z}) \propto \exp\left[\gamma \sum_{\boldsymbol{r} \in \mathcal{R}} \sum_{\boldsymbol{r}' \in \mathcal{V}(\boldsymbol{r})} \delta(z(\boldsymbol{r}) - z(\boldsymbol{r}'))\right] \\ m_{ik} \quad \text{follow a Markovian model along the index } i \end{array}$$



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Segmentation of a video sequence of images (with P. Brault)

$$\begin{aligned} g_i(\mathbf{r}) &= f_i(\mathbf{r}) + \epsilon_i(\mathbf{r}) \\ p(f_i(\mathbf{r})|z_i(\mathbf{r}) = k) &= \mathcal{N}(m_{ik}, \sigma_{ik}^2), \quad k = 1, \cdots, K \\ p(\underline{f}|\mathbf{z}) &= \prod_i p(f_i|\mathbf{z}_i) \\ p(\mathbf{z}) &\propto \exp\left[\gamma \sum_{\mathbf{r} \in \mathcal{R}} \sum_{\mathbf{r}' \in \mathcal{V}(\mathbf{r})} \delta(z(\mathbf{r}) - z(\mathbf{r}'))\right] \\ z_i(\mathbf{r}) \quad \text{follow a Markovian model along the index } i \end{aligned}$$



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Source separation: (with H. Snoussi & M. Ichir)

$$\begin{cases}
g_i(\mathbf{r}) = \sum_{j=1}^{N} A_{ij} f_j(\mathbf{r}) + \epsilon_i(\mathbf{r}) \\
p(f_j(\mathbf{r})|z_j(\mathbf{r}) = k) = \mathcal{N}(m_{j_k}, \sigma_{j_k}^2) \\
p(\mathbf{z}) \propto \exp\left[\gamma \sum_{\mathbf{r} \in \mathcal{R}} \sum_{\mathbf{r}' \in \mathcal{V}(\mathbf{r})} \delta(z(\mathbf{r}) - z(\mathbf{r}'))\right] \\
p(A_{ij}) = \mathcal{N}(A_{0ij}, \sigma_{0ij}^2)
\end{cases}$$



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Microwave imaging for breast cancer detection

(with L. Gharsally and B. Duchêne)



- Breast model built up from MRI scan [Zastrow et al. 2008]
- ▶ 64 sources
- 64 receivers
- ▶ 6 frequencies in the band 0.5 - 3 GHz

$\mathcal{D} = N_x \times N_y$	D1 $(\epsilon_1, \sigma_1(S/m))$	D2 $(\epsilon_2, \sigma_2(S/m))$	D3 $(\epsilon_3, \sigma_3(S/m))$	D4 $(\epsilon_3, \sigma_3(S/m))$
120×120	(10, 0.5)	(6.12, 0.11)	([2.46, 60.6], [0.01, 2.28])	(55.3, 1.57)

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Microwave imaging for breast cancer detection

CSI: Contrast Source Inversion, VBA: Variational Bayesian Approach, MGI: Independent Gaussian mixture, MGM: Gauss-Markov mixture



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Optical Diffraction Tomographic imaging

(with H. Ayasso and B. Duchêne)





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Optical Diffraction Tomographic imaging (with H. Ayasso and B. Duchêne)



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Acoustic source localization [Ning CHU et al]

Vehicle acoustic imaging at 2500Hz





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Acoustic source localization (Simulation)



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Conclusions

- Bayesian Inference for inverse problems
- Different prior modeling for signals and images: Separable, Markovian, without and with hidden variables
- Sprasity enforcing priors
- Gauss-Markov-Potts models for images incorporating hidden regions and contours
- Two main Bayesian computation tools: MCMC and VBA
- Application in different CT (X ray, Microwaves, PET, SPECT)

Current Projects and Perspectives :

- Efficient implementation in 2D and 3D cases
- Evaluation of performances and comparison between MCMC and VBA methods
- Application to other linear and non linear inverse problems: (PET, SPECT or ultrasound and microwave imaging)

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Current Applications and Perspectives

We use these models for inverse problems in different signal and image processing applications such as:

- Period estimation in biological time series
- ► Signal deconvolution in Proteomic and molecular imaging
- X ray Computed Tomography
- Diffraction Optical Tomography
- Microwave Imaging, Acoustic imaging and sources localization
- Synthetic Aperture Radar (SAR) Imaging

Thanks to:

Graduated PhD students:

- 1. C. Cai (2013: Multispectral X ray Tomography)
- 2. N. Chu (2013: Acoustic sources localization)
- 3. Th. Boulay (2013: Non Cooperative Radar Target Recognition)
- 4. R. Prenon (2013: Proteomic and Masse Spectrometry)
- 5. Sh. Zhu (2012: SAR Imaging)
- 6. D. Fall (2012: Emission Positón Tomography, Non Parametric Bayesian)
- 7. D. Pougaza (2011: Copula and Tomography)
- 8. H. Ayasso (2010: Optical Tomography, Variational Bayes)
- 9. S. Fékih-Salem (2009: 3D X ray Tomography)
- 10. N. Bali (2007: Hyperspectral imaging)
- 11. O. Féron (2006: Microwave imaging)
- 12. F. Humblot (2005: Super-resolution)
- 13. M. Ichir (2005: Image separation in Wavelet domain)
- 14. P. Brault (2005: Video segmentation using Wavelet domain)
- 15. H. Snoussi (2003: Sources separation)
- 16. Ch. Soussen (2000: Geometrical Tomography)
- 17. G. Montémont (2000: Detectors, Filtering)
- 18. H. Carfantan (1998: Microwave imaging)
- 19. S. Gautier (1996: Gamma ray imaging for NDT)
- 20. M. Nikolova (1994: Piecewise Gaussian models and GNC)
- 21. D. Prémel (1992: Eddy current imaging)

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- L. Gharsali (Microwave imaging for Cancer detection)
- M. Dumitru (Multivariate time series analysis for biological signals)
- S. AlAli (Electrical imaging of CO2 stocking under the earth)

Master students:

- A. Cai (Non-circular X ray Tomography)
- F. Fuc (Multi component signal analysis for biology applications)

Post-Docs:

- J. Lapuyade (2011: Dimentionality Reduction and multivariate analysis)
- S. Su (2006: Color image separation)
- A. Mohammadpour (2004: HyperSpectral image segmentation)

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- E. Barat (CEA-LIST) (Positon Emission Tomography, Non Parametric Bayesian)
- C. Comtat (SHFJ, CEA) (PET, Spatio-Temporal Brain activity)
- J. Picheral (SSE, Supélec) (Acoustic sources localization)
- D. Blacodon (ONERA) (Acoustic sources separation)
- J. Lagoutte (Thales Air Systems) (Non Cooperative Radar Target Recognition)
- P. Grangeat (LETI, CEA, Grenoble) (Proteomic and Masse Spectrometry)
- F. Lévi (CNRS-INSERM, Hopital Paul Brousse) (Biological rythms and Chronotherapy of Cancer)

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Current PhD's and projects

PhD's:

- 1. Microwave imaging: PhD Leila Gharsalli (co-supervising B. Duchêne)
- 2. Multivariate and multicomponents biological data processing: PhD Mircea Dumitru (co-supervising F. Lévi), ERASYSBIO
- 3. ANR: HONTOMIN, PhD Safa AlAli, (CO2 stock supervising using electrical imaging) (B. Duchne & G. Perruson)
- 4. New methods for reducing dose in Computed Tomography, PhD Li Wang (N. Gac)
- 5. Information fusion for radar target recognition, starting PhD, May Abou Chahine, Thales Systèmes Aéroports

Post-docs

- 1. ANR: SURMITO (Optical imaging), S. Mehrab, (B. Duchêne)
- 2. 3D Tomography (SAFRAN), Th. Boulay, (N. Gac)